



Canada Border
Services Agency

Agence des services
frontaliers du Canada



Calibrated (Probabilistic) Confidence Scoring for Biometric Identification

**Goal: from scores to probabilities
(0, .5, .5) → (80%, 10%, 10%)**

IBPC Conference (NIST), March 2-4, 2010

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Canada

CBSA - a prime user of Iris biometrics

Why iris ? – Easily accepted by public, touch-less / non-intrusive

Today: for collaborative user-engaged identification of pre-approved travellers in structured/overt environment (NEXUS)

Tomorrow: for fully-automated stand-off (on-the-fly) identification of Good and Bad people as they cross the border ?(3 persons crossing / sec)

Recent RFI examination (Feb 2009-Aug 2009) exposed the problems even with Today's systems/data

With Tomorrow's stand-off systems, these problems will be even more significant!

Gorodnichy, D. O. *“Evolution and evaluation of biometric systems”* IEEE Symposium: Computational Intelligence for Security and Defence Applications, Ottawa June 2009

Gorodnichy, D. O. *“Multi-order analysis framework for comprehensive biometric performance evaluation”*, SPIE Conf. on Defense, Security, and Sensing. Orlando, April 2010

Problems exposed through RFI



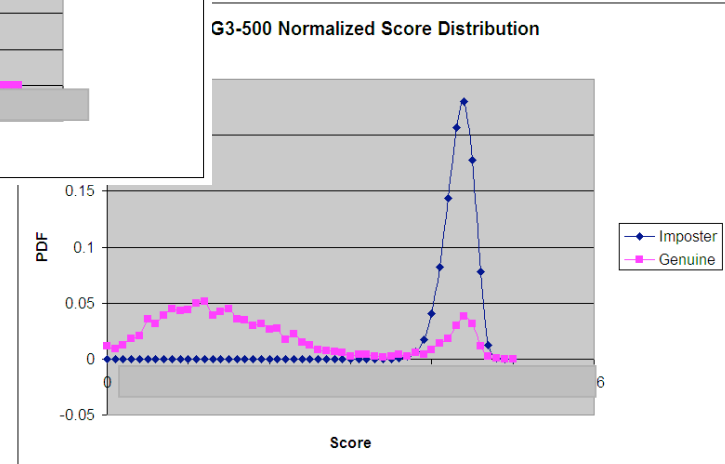
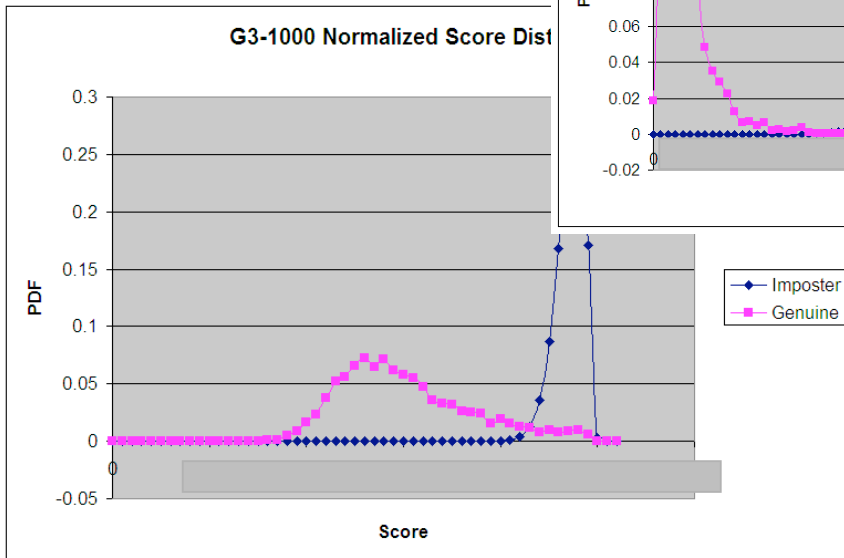
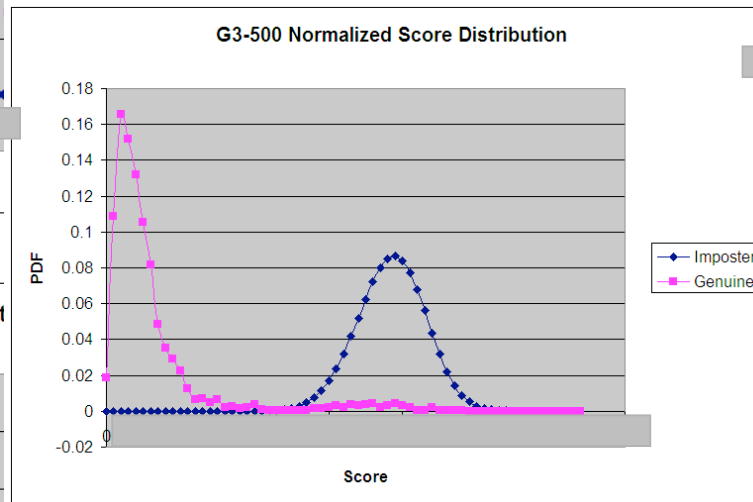
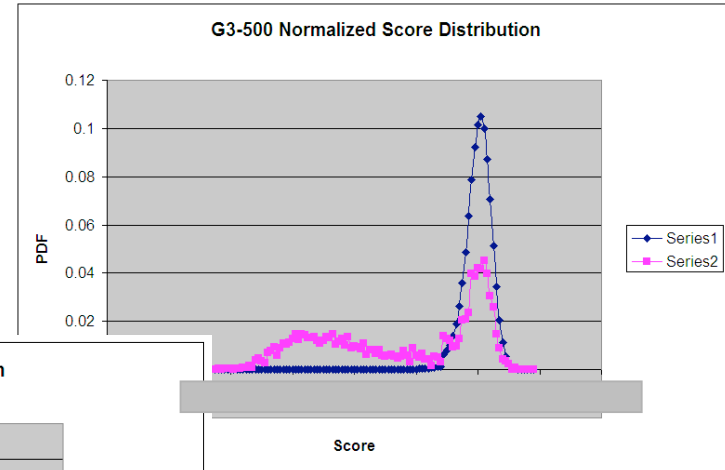
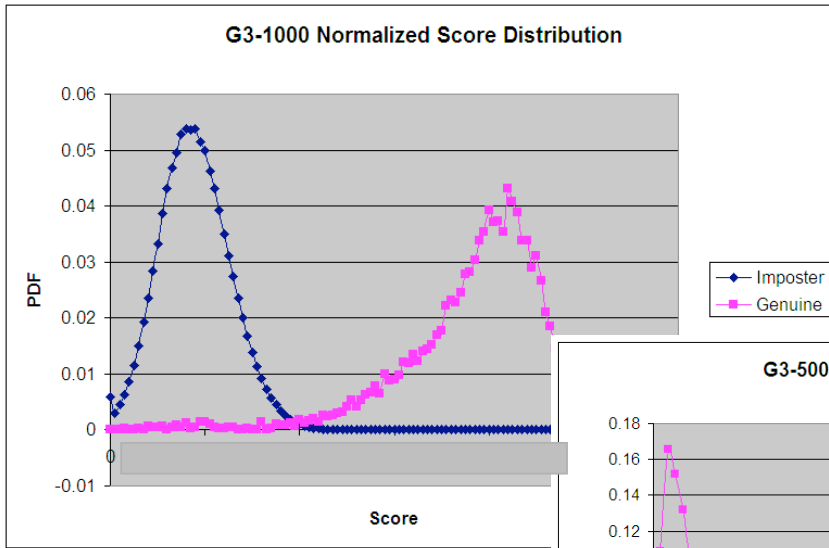
(With over 20.000.000 CBSA iris data, several state-of-art products, and over 6 months of coding and collecting/analyzing results)

1. There exist many (>5) matching algorithms now
 - All produce single scores output only (no confidence)!
 - Binomial nature of Imposter distributions
 - Binomial nature of Genuine distribution ? - with no noise
2. High FNMR (False Rejects, False Non-Match Rate)
3. High FTA (Failure To Acquire)
4. Despite many vendor/publications claims, systems often have :
 - 1) more than one match below the threshold,
 - 2) two or more close matching scores

There is a need therefore to assign *Confidence value to output!*



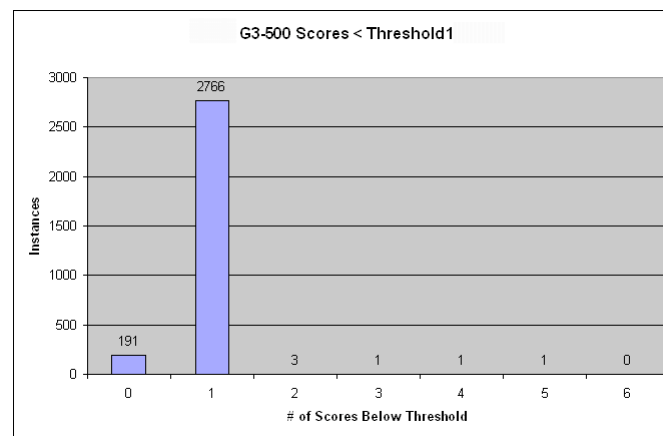
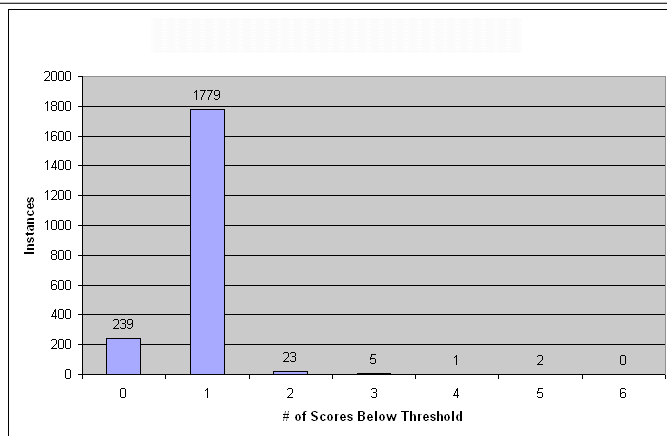
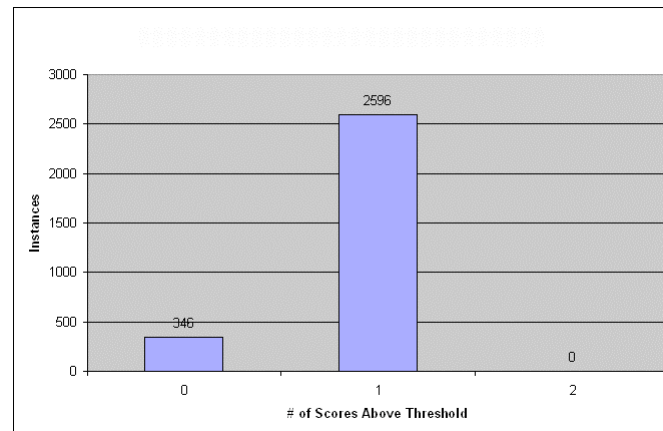
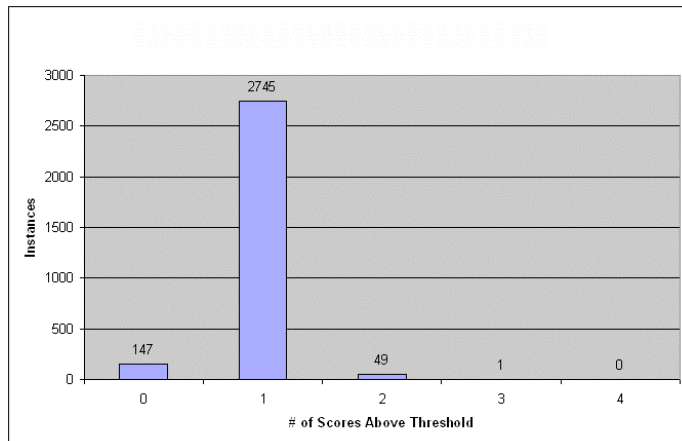
Anonymized score distributions



Anonymized stats

Using Multi-order score analysis [Gor09,10], Order 3 have shown that:

Many systems may improve FTA, FNMR, DET (match/non-match tradeoff) at the cost of allowing more than one score below a threshold

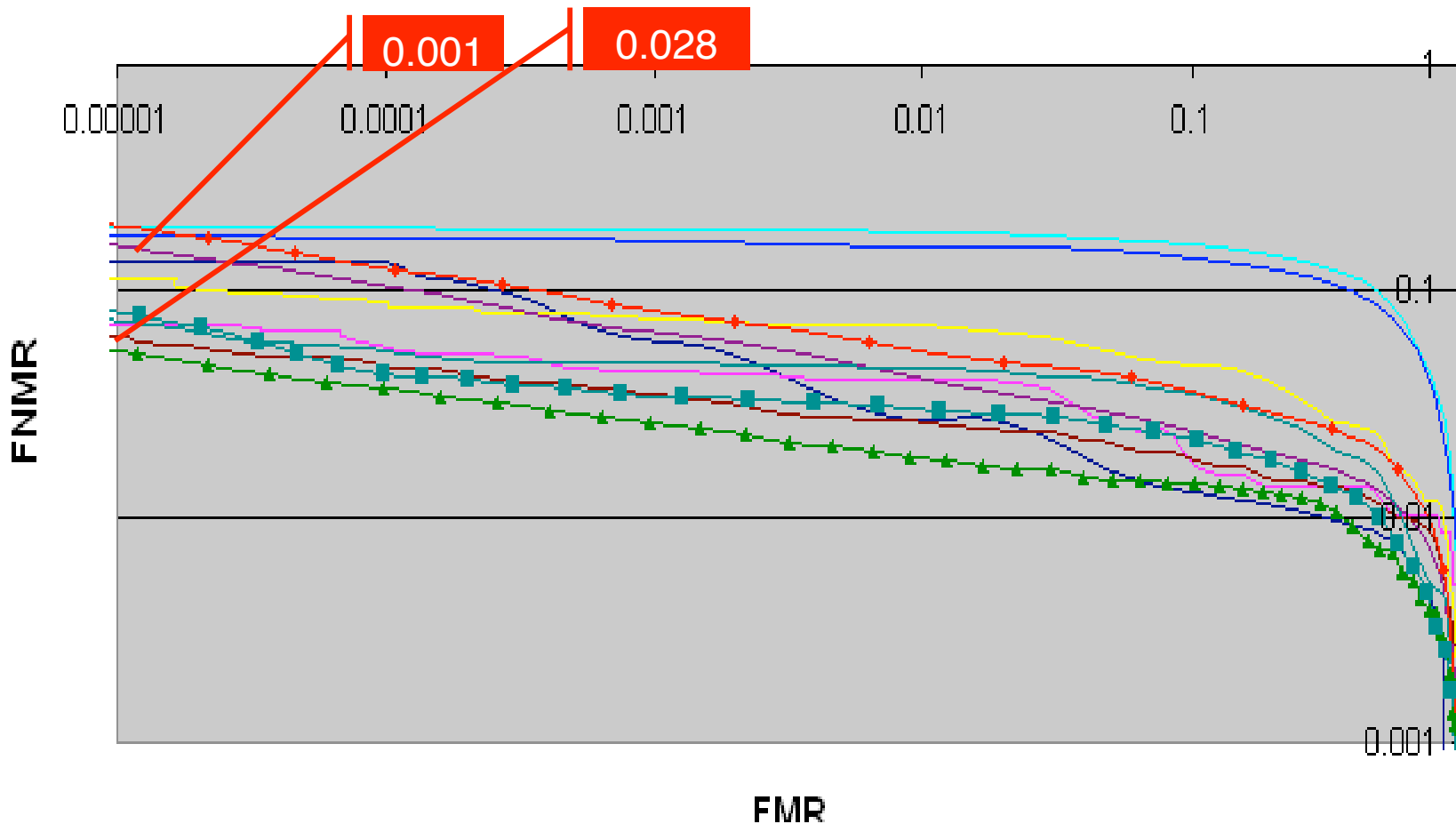


(With 500 enrolled travelers, each having 6 passage images)

Trade-off Curves with FCR



DEFINITION [Gor10]: **Failure of Confidence Rate (FCR)** – the rate of incidences in which there are more than one match below threshold



Goal: assign confidences to decisions

Given: Person X arrives at the kiosk and produces n scores:

n-tuple $S = (s_1, s_2, \dots, s_n)$, $s_i = \text{HD}(X, x_i)$

Find: Sequence of calibrated confidence scores:

the probability vector $C = (c_1, c_2, \dots, c_n)$, $c_i = P(\{X = x_i\} | S)$

How: as in probabilistic weather forecasting [DeGroot1983]

1. Make use of (assume) binomial nature of Genuine and Imposter score distributions [Daugman1993,2004]:

- $G \sim \text{Binom}(m', u')$, with $u' = 0.11$, $d' = 0.065$ ($m' \approx 115$).
- $I \sim \text{Binom}(m, u)$, with $u = 0.5$, $m = 249$ ($d \approx 0.03$)
- $P(\text{HD}=k/m) = \binom{k}{m} u^k (1-u)^{m-k}$

2. Bayes's Theorem for $c_i = P(\{X = x_i\} | S) =$

$$= P(\{X = x_i\} \wedge S) = P(\{X = x_i\} \wedge S) / P(S) = \dots$$

3. $P(\{X = x_i\} \wedge S) = \dots$

Simple example to illustrate



Enrolled: three individuals $\{x_1, x_2, x_3\}$, six bits in iris string.

- Thus, $n = 3$, $m = m' = 6$.
- $G = \text{Binom}(m', u')$, $I = \text{Binom}(m, u)$ with $u' = 1/3$ and $u = 1/2$.
- $x_1 = [0, 1, 0, 1, 0, 1]$, $x_2 = [1, 0, 0, 1, 1, 1]$, $x_3 = [1, 0, 1, 1, 0, 1]$

New person: $X = [0, 1, 0, 1, 0, 1]$.

- Matching scores $S = (0, 0.5, 0.5)$. Decision scores: $(1, 0, 0)$.

Using the theorem (for $q=0$ and $P_1=P_2=P_3$), we obtain:

- confidence scores $C = (0.8, 0.1, 0.1)$.

How to apply to real system?

- Vendor should provide: m', u', m, u
- User knows: P_i, q (a-priory probabilities of each person / imposter)

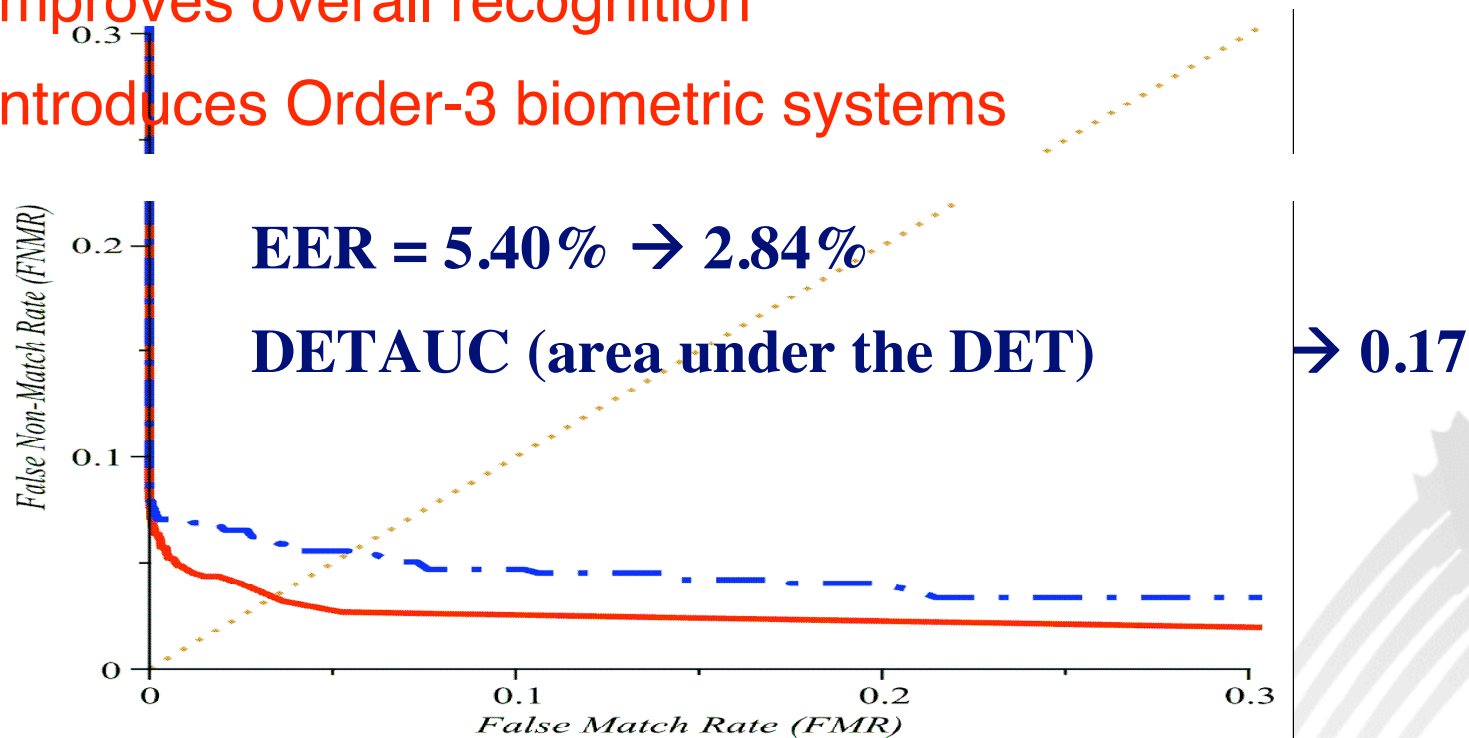


Applied to real system



Proposed probabilistic score calibration can be added to any system at little computation cost as post-processing filter:

- Provides more meaningful output - for risk mitigating procedures
- Improves overall recognition
- Introduces Order-3 biometric systems



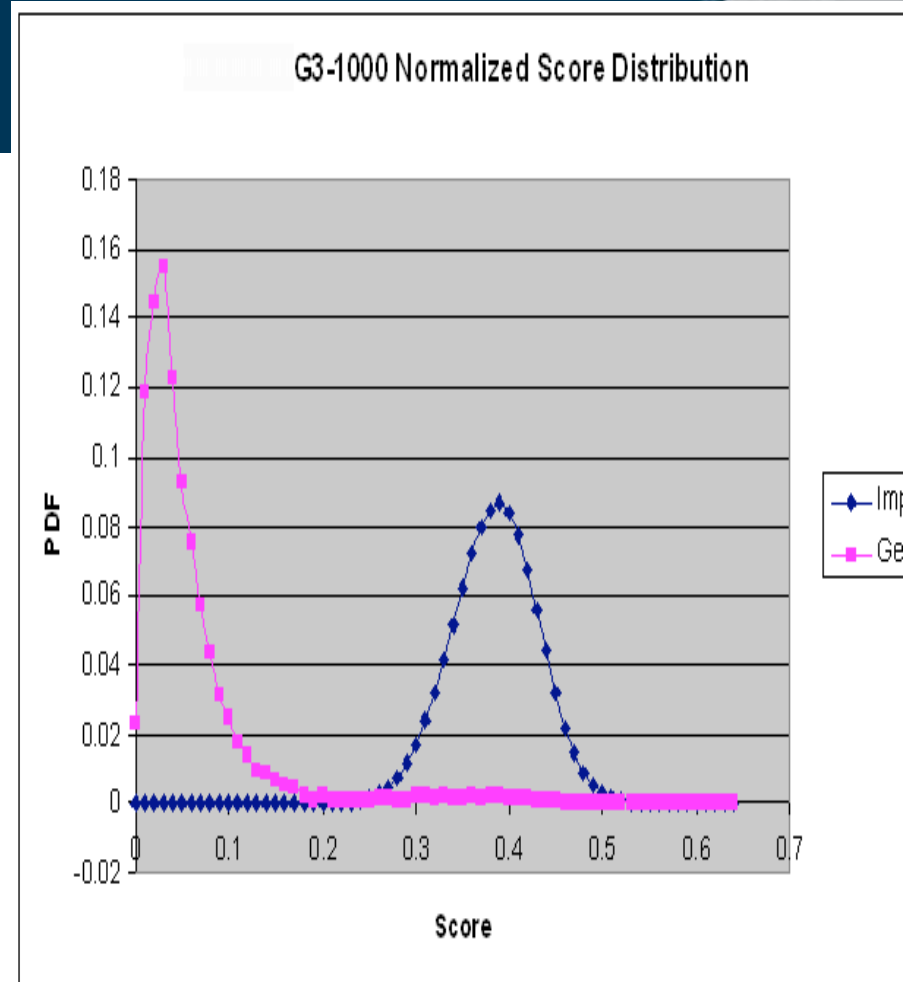


Appendices



Iris biometrics

- Image converted to 2048 binary digits {0,1}
 - only small subsets of bits are mutually independent [1].
- Impostor HD scores follow binomial distribution:
 $I \sim \text{Binom}(m, u)$,
 $m = 249$ and $u = 0.5$.
- The variable m represents the degrees-of-freedom and is a function of the mean u and the standard deviation d :
 $m = u(1 - u) / d^2$



- Genuine HD scores [2]:
 $G \sim \text{Binom}(m', u')$ with
 $u' = 0.11$, $d' = 0.065$

Main theorem and proof:



Theorem 3.1 Let G be the set of genuine matching scores, and I be the set of impostor matching scores. Suppose $G \sim \text{Binom}(\hat{m}, \hat{u})$ and $I \sim \text{Binom}(m, u)$. Let $p_i = P(X = x_i)$ and $q = 1 - \sum_{i=1}^n p_i$. Let $S = (s_1, s_2, \dots, s_n)$ be the n -tuple of matching scores produced by person X . Then for each $1 \leq i \leq n$, we have

$$c_i = P(X = x_i | S) = \frac{p_i z_i}{\sum_{i=1}^n p_i z_i + q \cdot \frac{(1-u)^m}{(1-\hat{u})^{\hat{m}}}}, \quad \text{where } z_i = \frac{\binom{\hat{m}}{\hat{m}s_i}}{\binom{m}{ms_i}} \cdot \left(\frac{\hat{u}^{\hat{m}}(1-u)^m}{u^m(1-\hat{u})^{\hat{m}}} \right)^{s_i}.$$

Proof: For each $1 \leq i \leq n$, define $r_i = P(\{X = x_i\} \wedge S)$. Also define $r_{imp} = P(\{X \notin \{x_1, x_2, \dots, x_n\}\} \wedge S)$.

By definition, $r_{imp} = P(S) - \sum_{i=1}^n r_i$. By Bayes' Theorem, we have

$$c_i = P(\{X = x_i\} | S) = \frac{P(\{X = x_i\} \wedge S)}{P(S)} = \frac{r_i}{r_1 + r_2 + \dots + r_n + r_{imp}}.$$

To calculate $r_i = P(\{X = x_i\} \wedge S)$, we multiply the probabilities of the following $n + 1$ independent events: it is x_i who comes to the kiosk; the genuine matching score $HD(X, x_i)$ is s_i ; and the impostor matching score $HD(X, x_j)$ is s_j for all $1 \leq j \leq n$ with $j \neq i$.

Since $G \sim \text{Binom}(\hat{m}, \hat{u})$, there are \hat{m} degrees-of-freedom, and the probability that any of these \hat{m} bits differ is \hat{u} . So if $HD(X, x_i) = s_i$, then $\hat{m}s_i$ of the \hat{m} bits differ. We derive the analogous result for the impostor distribution $I \sim \text{Binom}(m, u)$, for all $1 \leq j \leq n$ with $j \neq i$. Therefore, we have

$$r_i = p_i \binom{\hat{m}}{\hat{m}s_i} \hat{u}^{\hat{m}s_i} (1-\hat{u})^{\hat{m}-\hat{m}s_i} \cdot \prod_{j=1, j \neq i}^n \binom{m}{ms_j} u^{ms_j} (1-u)^{m-ms_j}$$

Details of our simple example



Because $m = m' = 6$, and $u = 1 - u' = 1/2$, $2 * u' = 1 - u' = 2/3$ many things get cancelled out ...

$$Z_i(S_i) = \binom{6}{6 * S_i} / \binom{6}{6 * S_i} * \left(\left(\frac{1}{3} \right)^6 * \left(\frac{1}{2} \right)^6 \right) / \left(\left(\frac{1}{2} \right)^6 * \left(\frac{2}{3} \right)^6 \right) \wedge S_i = \left(\frac{1}{2} \right)^6 \wedge S_i = \left(\frac{1}{2} \right)^{6 * S_i}$$

For $S_2 = S_3 = 0.5$, we have: $Z_2 = Z_3 = \left(\frac{1}{2} \right)^3 = 1/8$.

For $S_1 = 0$, $Z_1 = 1$

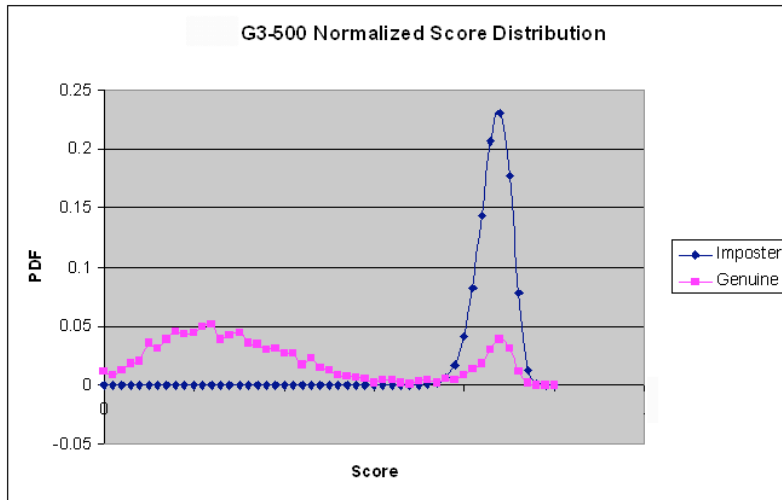
Then $C_i = (Z_i) / (\text{SUM } Z_i) = Z_i / (1/8 + 1/8 + 1)$

and $C_2 = 4/5 * (1/8) = 1/10$, $C_1 = 8/10$

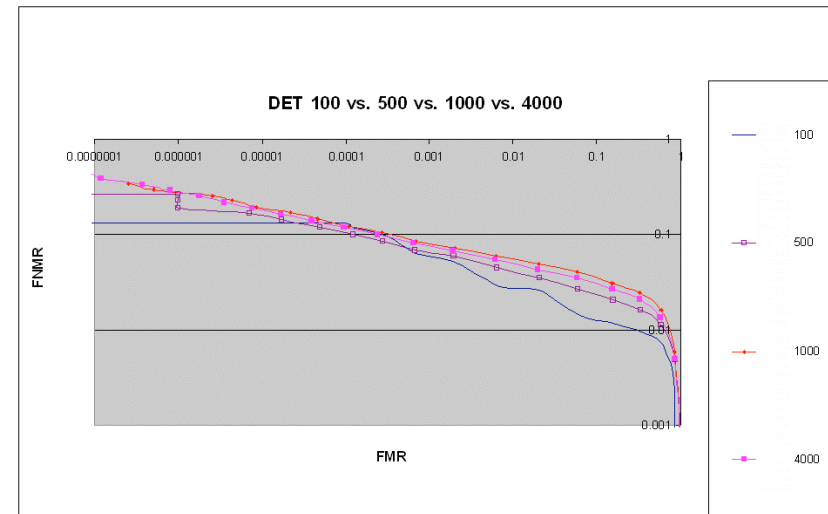


Multi-order performance evaluation

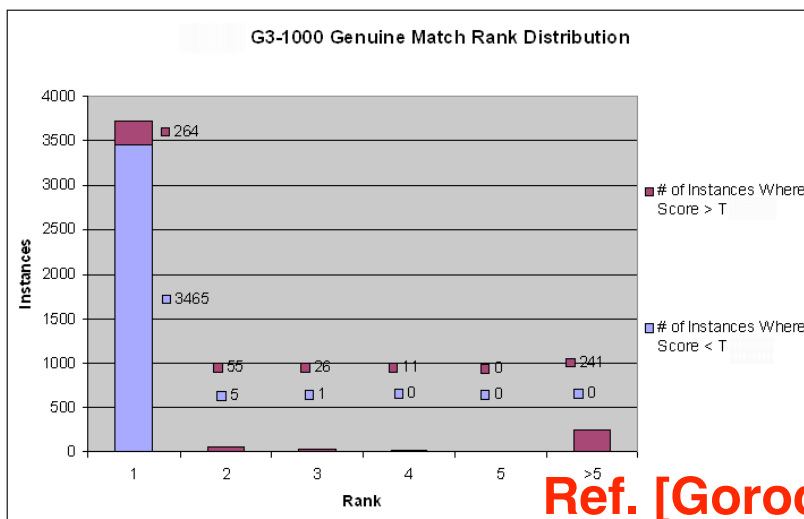
Order 0:



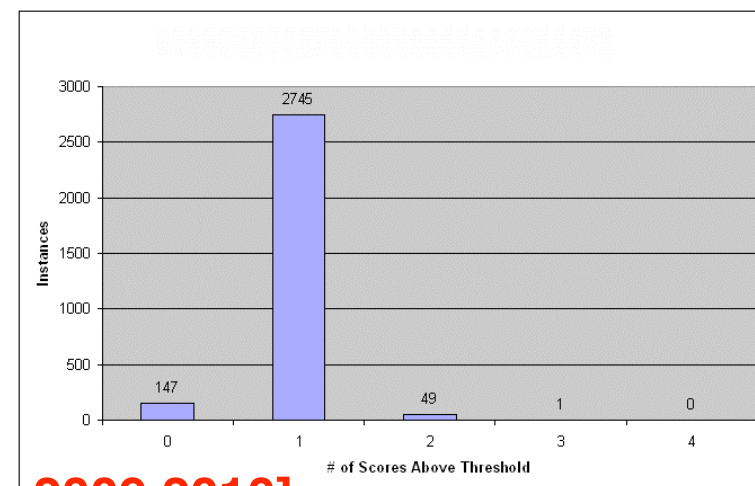
Order 1:



Order 2:



Order 3:



Ref. [Gorodnichy2009,2010]

Multi-order score analysis



Order 1 (Traditional):

- Examine single-scores to report trade-off (FMR/FNMR) curves

Order 2:

- Examine all scores to report the best (smallest) score

Order 3:

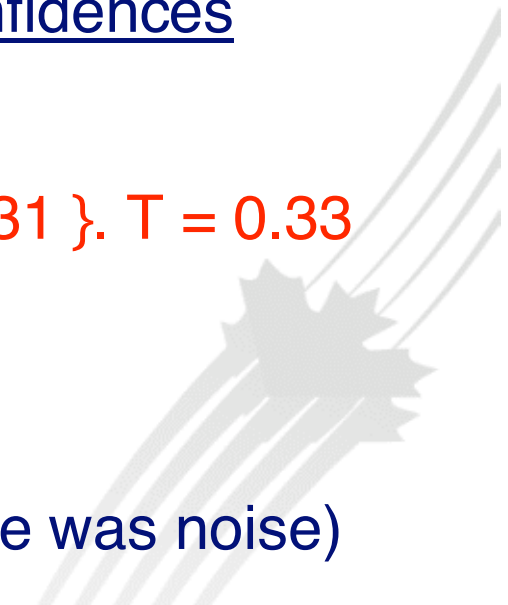
- Examine all scores relationship to report Confidences

Five-score example: { 0.51, 0.32, 0.47, 0.34, 0.31 }. $T = 0.33$

- Order 1 → 0.32

- Order 2 → 0.31

- But in reality it could have been 0.34 ! (if there was noise)



References

- Daugman, J. (1993). High confidence visual recognition of persons by a test of statistical independence. *IEEE Transactions on Pattern Analysis and Machine Intelligence*,
- Daugman, J. (2004). How iris recognition works. *IEEE Transactions on Circuits and Systems for Video Technology*, 14(1) 21-30.
- DeGroot, M. H. and Fienberg, S. E. (1983), The comparison and evaluation of probability forecasters, *Statistician*, 12, 12-22.
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